

How can I use the Imatrix subcommand to understand a three-way interaction in ANOVA?

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The Imatrix subcommand is a useful tool for understanding a three-way interaction in ANOVA (Analysis of Variance). This subcommand allows for the creation of a matrix that displays the means and standard errors for each combination of the three factors in the ANOVA model. By examining this matrix, one can gain a better understanding of how the three factors interact with each other and their overall effect on the dependent variable. This can aid in interpreting the results of the ANOVA and identifying significant interactions. Overall, the Imatrix subcommand can be a valuable tool in analyzing and interpreting complex three-way interactions in ANOVA.

How can I use the Imatrix subcommand to understand a three-way interaction in ANOVA? | SPSS FAQ

You can use multiple Imatrix subcommands to explore the interaction of three categorical variables in ANOVA. If you are not familiar with three-way interactions in ANOVA, please see our general FAQ on understanding three-way interactions in ANOVA. In short, a three-way interaction means that there is a two-way interaction that varies across levels of a third variable. Say, for example, that a b*c interaction differs across various levels of factor a.

One way of analyzing the three-way interaction is through the use of tests of simple main-effects, e.g.,

the effect of one variable (or set of variables) across the levels of another variable.

We will use a small artificial dataset called threeway that has a statistically significant three-way interaction to illustrate the process. In our example data set, variables a, b and

c

are categorical. The techniques shown on this page can be generalized to

situations in which one or more variables are continuous, but the more continuous variables that are involved in the interaction, the more complicated things get.

The results (shown below) indicate that the b*c interaction is statistically significant at a=1 but not at a=2. Because of this, the second two

lmatrix subcommands are needed; these show the effect of c at a=1 at both levels of b.

After we look at the results, we will look at the coding

used.

glm y by a b c

/design = a b c a*b a*c b*c a*b*c

**/lmatrix 'b*c at a=1' b*c 1 0 -1 -1 0 1 a*b*c 1 0 -1 -1 0 1 0 0
0 0 0 0;**

b*c 0 1 -1 0 -1 1 a*b*c 0 1 -1 0 -1 1 0 0 0 0 0 0

**/lmatrix 'b*c at a=2' b*c 1 0 -1 -1 0 1 a*b*c 0 0 0 0 0 0 1 0
-1 -1 0 1;**

b*c 0 1 -1 0 -1 1 a*b*c 0 0 0 0 0 0 0 1 -1 0 -1 1

**/lmatrix 'c at a=1 & b=1' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 1 0 -1
0 0 0 a*b*c 1 0 -1 0 0 0 0 0 0 0 0 0;**

**c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 1 -1 0 0 0 a*b*c 0 1 -1 0 0 0
0 0 0 0 0**

**/lmatrix 'c at a=1 & b=2' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 0 0 0
1 0 -1 a*b*c 0 0 0 1 0 -1 0 0 0 0 0 0;**

**c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 0 0 0 1 -1 a*b*c 0 0 0 0 1 -1 0
0 0 0 0 0.**

Between-Subjects Factors

		N
a	1.00	12
	2.00	12
b	1.00	12
	2.00	12
c	1.00	8
	2.00	8
	3.00	8

Tests of Between-Subjects Effects

Dependent Variable: y

Source	Type III Sum of Squares	df	Mean Square	F	Sig.
Corrected Model	497.833 ^a	11	45.258	33.943	.000
Intercept	5460.167	1	5460.167	4095.125	.000
a	150.000	1	150.000	112.500	.000
b	.667	1	.667	.500	.493
c	127.583	2	63.792	47.844	.000
a * b	160.167	1	160.167	120.125	.000
a * c	18.250	2	9.125	6.844	.010
b * c	22.583	2	11.292	8.469	.005
a * b * c	18.583	2	9.292	6.969	.010
Error	16.000	12	1.333		
Total	5974.000	24			
Corrected Total	513.833	23			

a. R Squared = .969 (Adjusted R Squared = .940)

Custom Hypothesis Tests Index

1	Contrast Coefficients (L' Matrix) Transformation Coefficients (M Matrix) Contrast Results (K Matrix)	LMATRIX Subcommand 1: b*c at a = 1 Identity Matrix Zero Matrix
2	Contrast Coefficients (L' Matrix) Transformation Coefficients (M Matrix) Contrast Results (K Matrix)	LMATRIX Subcommand 2: b*c at a=2 Identity Matrix Zero Matrix
3	Contrast Coefficients (L' Matrix) Transformation Coefficients (M Matrix) Contrast Results (K Matrix)	LMATRIX Subcommand 3: a = 1 and b = 1 comparing c Identity Matrix Zero Matrix
4	Contrast Coefficients (L' Matrix) Transformation Coefficients (M Matrix) Contrast Results (K Matrix)	LMATRIX Subcommand 4: c at a=1 & b=2 Identity Matrix Zero Matrix

Custom Hypothesis Tests #1

Contrast Results (K Matrix)^a

Contrast		Dependent Variable	
		y	
L1	Contrast Estimate	-9.000	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	-9.000	
	Std. Error	1.633	
	Sig.	.000	
	95% Confidence Interval for Difference	Lower Bound	-12.558
	Upper Bound	-5.442	
L2	Contrast Estimate	-5.000	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	-5.000	
	Std. Error	1.633	
	Sig.	.010	
	95% Confidence Interval for Difference	Lower Bound	-8.558
	Upper Bound	-1.442	

a. Based on the user-specified contrast coefficients (L) matrix: b*c at a = 1

Test Results

Dependent Variable: y

Source	Sum of Squares	df	Mean Square	F	Sig.
Contrast	40.667	2	20.333	15.250	.001
Error	16.000	12	1.333		

Custom Hypothesis Tests #2

Contrast Results (K Matrix)^a

Contrast		Dependent Variable	
		y	
L1	Contrast Estimate	-.500	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	-.500	
	Std. Error	1.633	
	Sig.	.765	
	95% Confidence Interval for Difference	Lower Bound Upper Bound	-4.058 3.058
	L2	Contrast Estimate	.500
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	.500	
	Std. Error	1.633	
	Sig.	.765	
	95% Confidence Interval for Difference	Lower Bound Upper Bound	-3.058 4.058

a. Based on the user-specified contrast coefficients (L) matrix: b*c at a=2

Test Results

Dependent Variable: y

Source	Sum of Squares	df	Mean Square	F	Sig.
Contrast	.500	2	.250	.188	.831
Error	16.000	12	1.333		

Custom Hypothesis Tests #3

Contrast Results (K Matrix)^a

Contrast		Dependent Variable	
		y	
L1	Contrast Estimate	-8.000	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	-8.000	
	Std. Error	1.155	
	Sig.	.000	
	95% Confidence Interval for Difference	Lower Bound Upper Bound	-10.516 -5.484
	L2	Contrast Estimate	-4.000
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	-4.000	
	Std. Error	1.155	
	Sig.	.005	
	95% Confidence Interval for Difference	Lower Bound Upper Bound	-6.516 -1.484

a. Based on the user specified contrast coefficients (L) matrix a = 1 and b = 1 comparing c

Test Results

Dependent Variable: y

Source	Sum of Squares	df	Mean Square	F	Sig.
Contrast	64.000	2	32.000	24.000	.000
Error	16.000	12	1.333		

Custom Hypothesis Tests #4

Contrast Results (K Matrix)^a

Contrast		Dependent Variable	
		y	
L1	Contrast Estimate	1.000	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	1.000	
	Std. Error	1.155	
	Sig.	.403	
	95% Confidence Interval for Difference	Lower Bound	-1.516
		Upper Bound	3.516
L2	Contrast Estimate	1.000	
	Hypothesized Value	0	
	Difference (Estimate - Hypothesized)	1.000	
	Std. Error	1.155	
	Sig.	.403	
	95% Confidence Interval for Difference	Lower Bound	-1.516
		Upper Bound	3.516

a. Based on the user-specified contrast coefficients (L) matrix: c at a=1 & b=2

Test Results

Dependent Variable: y

Source	Sum of Squares	df	Mean Square	F	Sig.
Contrast	1.333	2	.667	.500	.619
Error	16.000	12	1.333		

In the first Imatrix subcommand, we are interested in the b*c

interaction at a=1. The b*c interaction has 2 degrees of freedom (

$(2-1)*(3-1) = 2$). To indicate this, we use a semicolon to separate the

two parts. Also, because we have included the two-way interaction, we also need to include the three-way interaction. In the second lmatrix subcommand, we are looking at the b*c interaction at a=2. Realistically, we wouldn't know to include the third and fourth lmatrix subcommands until we had run the first two and seen the results. To save space, we have included these two lmatrix subcommands, which investigate c at a=1 and both levels of b.

Let's look a little closer at the coding of the variables on the lmatrix subcommands. First, we need to remember that the variable a has two levels, b has two levels, and c has three levels. The coding (which is effect coding) is for each cell produced by the crossing of the categorical predictor variables. This is perhaps

best understood as the "differences of differences" approach. (For more information, please see

Multiple

Regression: Testing and Interpreting Interactions by Leona S. Aiken and Steven G. West).

```
glm y by a b c
```

```
/design = a b c a*b a*c b*c a*b*c
```

```
/lmatrix 'b*c at a=1' b*c 1 0 -1 -1 0 1 a*b*c 1 0 -1 -1 0 1 0 0
0 0 0 0;
```

```
b*c 0 1 -1 0 -1 1 a*b*c 0 1 -1 0 -1 1 0 0 0 0 0 0
```

```
/lmatrix 'b*c at a=2' b*c 1 0 -1 -1 0 1 a*b*c 0 0 0 0 0 0 0 1 0
-1 -1 0 1;
```

```
b*c 0 1 -1 0 -1 1 a*b*c 0 0 0 0 0 0 0 1 -1 0 -1 1
```

```
/lmatrix 'c at a=1 & b=1' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 1 0 -1
0 0 0 a*b*c 1 0 -1 0 0 0 0 0 0 0 0 0 0 0 0;
```

```
c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 1 -1 0 0 0 a*b*c 0 1 -1 0 0 0 0
0 0 0 0 0
```

```
/lmatrix 'c at a=1 & b=2' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 0 0 0
1 0 -1 a*b*c 0 0 0 1 0 -1 0 0 0 0 0 0 0 0;
```

**c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 0 0 0 1 -1 a*b*c 0 0 0 0 1 -1 0
0 0 0 0 0.**

The first Imatrix subcommand

Let's take the first line of the first Imatrix subcommand as an

example. We have the b*c interaction at a=1, and we are comparing c1 to

c3. In other words, c3 is our reference group. Picking c3 as our

reference group is somewhat arbitrary; we could have used c1 or c2. The

"differences of differences" approach means that we are going to take the

difference of c1 and c3 at b=1, and the difference of c1 and c3 at b=2, and then

take the difference of those two differences. In the table below, we have

six cells (because 2 levels of b times 3 levels of c equals

6). We have called the cells msubscript, so that we can do

some symbolic math.

a=1

	c1	c2	c3
b=1	m11	m12	m13
b=2	m21	m22	m23

$$(m11 - m13) - (m21 - m23)$$

$$(1 \ 0 \ -1) - (1 \ 0 \ -1) = 1 \ 0 \ -1 \ -1 \ 0 \ 1$$

Notice that 1 0 -1 -1 0 1 are the first six entries in the first line of the first Imatrix subcommand.

Now let's look at the second part, the a*b*c interaction. The first six numbers are for a=1, and the second six are for a=2. Because we are only looking at a=1 in this analysis, all of the values for a=2 are 0. The values for a=1 are the same as those for the b*c interaction.

Here is another way of thinking about the first line of the first Imatrix subcommand:

```
/lmatrix 'b*c at a=1' b*c 1 0 -1 -1 0 1 a*b*c 1 0 -1 -1 0 1 0 0
00 0 0;
```

Yellow: b=1, comparing c1 with c3

Orange: b=2, comparing c1 with c3

Green: a=1 and b=1, comparing c1 with c3

Blue: a=1 and b=2, comparing c1 with c3

Pink: a=2 and b=1, these are all 0s because we are looking only at a=1

Purple: a=2 and b=2, these are all 0s because we are looking only at a=1

The second line of the first Imatrix subcommand is very similar to the first, except that it is for c2 versus c3. So, we have

$$(m_{12} - m_{13}) - (m_{22} - m_{23})$$

$$(0 \ 1 \ -1) - (0 \ 1 \ -1) = 0 \ 1 \ -1 \ 0 \ -1 \ 1$$

```
/lmatrix 'b*c at a=1' b*c 1 0 -1 -1 0 1 a*b*c 1 0 -1 -1 0 1 0 0
0 0 0 0;
b*c 0 1 -1 0 -1 1 a*b*c 0 1 -1 0 -1 1 0 0 0 0 0
```

Yellow: b=1, comparing c2 with c3

Orange: b=2, comparing c2 with c3

Green: a=1 and b=1, comparing c2 with c3

Blue: a=1 and b=2, comparing c2 with c3

Pink: a=2 and b=1, these are all 0s because we are looking only at a=1

Purple: a=2 and b=2, these are all 0s because we are looking only at a=1

The second lmatrix subcommand

The second lmatrix subcommand looks at the b*c interaction at a=2.

It is the same as the first, except in the part for the a*b*c interaction. Here, the first six 0s are for a=1, which

we are not considering in this lmatrix subcommand.

The same coding

used in the first lmatrix subcommand is simply shifted to the a=2 part of the code.

The third lmatrix subcommand

```
/lmatrix 'c at a=1 & b=1' c 1 0 -1 a*c 1 0 -10 0 0 b*c 1 0 -10
0 0 a*b*c 1 0 -10 0 0 0 0 0 0;
c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 1 -1 0 0 0 a*b*c 0 1 -1 0 0 0
```

0 0 0 0 0

By now, the coding for **c**, the first part of the Imatrix subcommand, should be familiar. In this first line, we are comparing **c1** with **c3**.

a=1

	c1	c2	c3
b=1	m11	m12	m13
b=2	m21	m22	m23

(m11 - m13) - (m21 - m23)

(1 0 -1) - (1 0 -1) = 1 0 -1 -1 0 1

Red: comparing c1 with c3

Light blue: a=1, comparing c1 and c3

Dark green: a=2, these are all 0 because we are looking at a=1

Yellow: b=1, comparing c1 with c3

Orange: b=2, these are all 0 because we are looking at b=1

Light green: a=1, b=1, comparing c1 with c3

Dark blue: a=1, b=2, these are all 0 because we are

looking at b=1

Pink: a=2, b=1, these are all 0 because we are looking at a=1

Purple: a=2, b=2, these are all 0 because we are looking at a=1 and b=1

The second line of the third Imatrix subcommand is very similar to the first

line, except that it compares c2 to c3.

```
/Imatrix 'c at a=1 & b=1' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 1 0 -1
0 0 0 a*b*c 1 0 -1 0 0 0 0 0 0 0 0 0;
c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 1 -1 0 0 0 a*b*c 0 1 -1 0 0 0 0
00 0 0
```

Light blue: a=1, comparing c2 and c3

Dark green: a=2, these are all 0 because we are looking at a=1

Yellow: b=1, comparing c2 with c3

Orange: b=2, these are all 0 because we are looking at b=1

Light green: a=1, b=1, comparing c2 with c3

Dark blue: a=1, b=2, these are all 0 because we are looking at b=1

Pink: a=2, b=1, these are all 0 because we are looking at a=1

Purple: a=2, b=2, these are all 0 because we are looking at a=1 and b=1

The fourth Imatrix subcommand

The fourth Imatrix subcommand is the same as the third, except we are now looking at b=2. Hence, we have 0s for the b=1 part of the code and the comparisons of the different levels of c in the b=2 part of the code.

```
/Imatrix 'c at a=1 & b=2' c 1 0 -1 a*c 1 0 -1 0 0 0 b*c 0 0 0
1 0 -1 a*b*c 0 0 0 1 0 -1 0 0 0 0 0 0;
c 0 1 -1 a*c 0 1 -1 0 0 0 b*c 0 0 0 0 1 -1 a*b*c 0 0 0 0 1 -1 0
0 0 0 0 0.
```

Correcting for multiple tests

We should note that although a p-value is given for each of the four F-tests, it is not corrected for the multiple tests. There are at least four different methods of determining the critical value of tests of

simple main-effects.

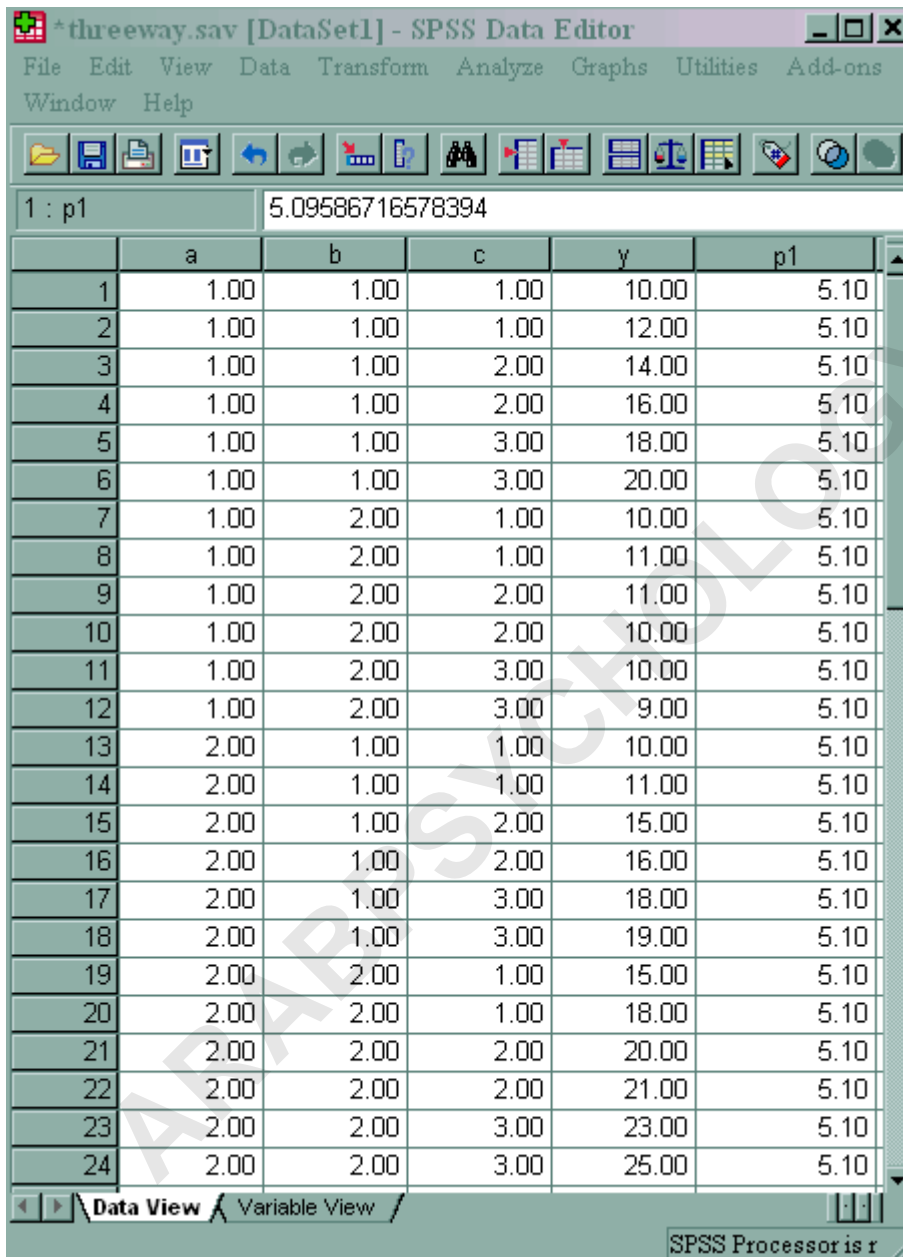
There is a method related to Dunn's multiple comparisons, a method attributed to Marascuilo and Levin, a method called the simultaneous test procedure (very conservative and related to the Scheffé post-hoc test) and a per family error rate method. We will demonstrate the per family error rate method, but you should look up the other methods in a good ANOVA book, such as Kirk (1995), to decide which approach is best for your situation.

Let's take the first two tests, comparing b*c at a=1 and at a=2 as an example. The values for the F-tests were 15.25 and .188, respectively.

We divide our alpha level, 0.05, by 2 because we are doing two tests of simple main-effects, so our new value of alpha is .025. The idf function requires us to provide 1 - alpha, so we have 1 - .025 = .975.

compute p1 = idf.f(.975, 2, 12).

exe.



The screenshot shows the SPSS Data Editor window for a file named 'threeaway.sav'. The window displays a data table with 24 rows and 6 columns. The columns are labeled 'a', 'b', 'c', 'y', and 'p1'. The 'p1' column contains the value 5.10 for all rows. The 'y' column contains values ranging from 9.00 to 25.00. The 'a', 'b', and 'c' columns contain values 1.00, 2.00, and 3.00, representing different levels of the factors. The status bar at the bottom indicates 'SPSS Processor is r'.

	a	b	c	y	p1
1	1.00	1.00	1.00	10.00	5.10
2	1.00	1.00	1.00	12.00	5.10
3	1.00	1.00	2.00	14.00	5.10
4	1.00	1.00	2.00	16.00	5.10
5	1.00	1.00	3.00	18.00	5.10
6	1.00	1.00	3.00	20.00	5.10
7	1.00	2.00	1.00	10.00	5.10
8	1.00	2.00	1.00	11.00	5.10
9	1.00	2.00	2.00	11.00	5.10
10	1.00	2.00	2.00	10.00	5.10
11	1.00	2.00	3.00	10.00	5.10
12	1.00	2.00	3.00	9.00	5.10
13	2.00	1.00	1.00	10.00	5.10
14	2.00	1.00	1.00	11.00	5.10
15	2.00	1.00	2.00	15.00	5.10
16	2.00	1.00	2.00	16.00	5.10
17	2.00	1.00	3.00	18.00	5.10
18	2.00	1.00	3.00	19.00	5.10
19	2.00	2.00	1.00	15.00	5.10
20	2.00	2.00	1.00	18.00	5.10
21	2.00	2.00	2.00	20.00	5.10
22	2.00	2.00	2.00	21.00	5.10
23	2.00	2.00	3.00	23.00	5.10
24	2.00	2.00	3.00	25.00	5.10

As you can see, p1 is approximately 5.10. This indicates that the b*c

interaction is statistically significant at $\alpha=1$ but not at $\alpha=2$.

References

Kirk, Roger E. (1995) Experimental Design: Procedures for the Behavioral Sciences, Third Edition. Monterey, California: Brooks/Cole Publishing.

Aiken, Leona S., and West, Stephen G. (1996) Multiple Regression: Testing and Interpreting Interactions. Thousand Oaks, California: Sage Publishing.