

“Can you explain the concept of random coefficient Poisson models in R?”

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RECOMMENDED CITATION

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The random coefficient Poisson model is a statistical method used to analyze count data where the count is assumed to follow a Poisson distribution. In this model, the expected value of the count is allowed to vary based on certain explanatory variables, known as random coefficients. These coefficients are assumed to follow a normal distribution, allowing for the incorporation of random effects into the model. In R, this concept can be implemented using the `glm()` function, with the addition of the random parameter to specify the random effects structure. This approach is useful for accounting for unobserved heterogeneity in the data and can provide more accurate predictions compared to traditional Poisson models.

Random Coefficient Poisson Models | R FAQ

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Please Note: The purpose of this page is to show how to use various data analysis commands.

It does not cover all aspects of the research process which researchers are expected to do. In particular, it does not cover data cleaning and checking, verification of assumptions, model diagnostics and potential follow-up analyses.

Examples of random coefficient poisson models

Poisson models are useful for count data. Examples could include number of traffic tickets an individual receives in a year, number of tumor sites in cancer

patients, or number of awards received by students. In each of these cases, we might expect most people to have very few, with a relatively small number of individuals having higher numbers.

Background on the Poisson distribution

Unlike the familiar Gaussian distribution which has two parameters (μ, σ^2), the Poisson distribution is described by a single parameter, (λ) that is both the mean and variance. That is, ($\lambda = E(x)$) and ($\lambda = \text{Var}(x) = E(x^2) - E(x)^2$).

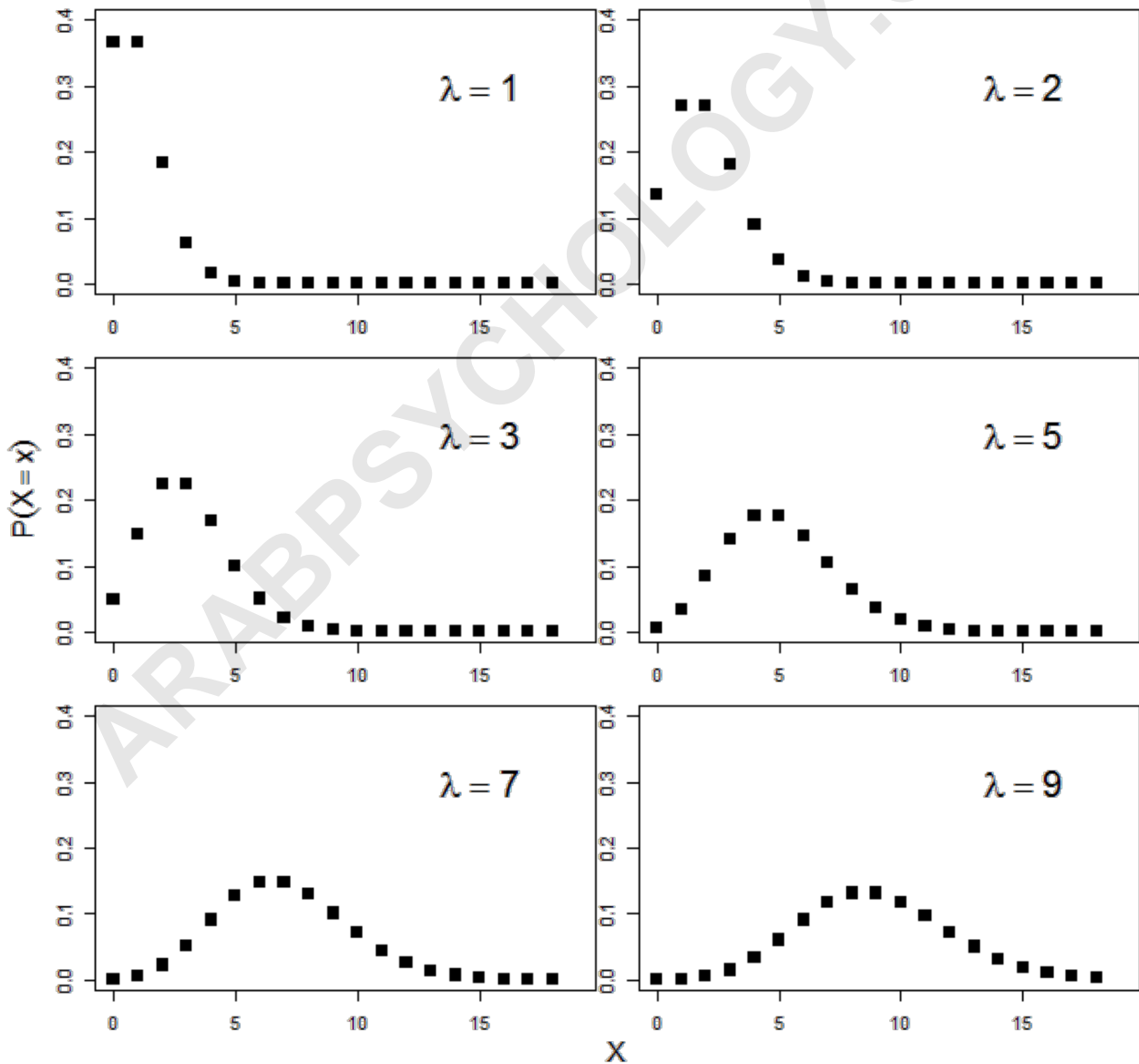
In many ways, this can be a strong assumption. In some cases, use of the negative binomial distribution which estimates a second scale parameter to allow for overdispersion. Here is what the Poisson distribution looks like for different values of (λ).

```
par(mfrow=c(3,2),mar=c(1,1,2,1),oma=c(4,4,2,0))for(iinc(1,2,3,5,7,9)){curve(dpois(x,lambda=i),from=0,to=18,n=19,type="p",pch=15,cex=1.5,xlim=c(0,19),ylim=c(0,0.4),xlab="",ylab="")text(x=15,y=0.3,substitute(lambda == x,list(x=i)),cex=2)}mtext("Probability Mass Function for
```

Poisson

Distribution",line=0.5,outer=TRUE,cex=1.2)mtext(expression(P(X == x)),side=2,line=1.5,outer=TRUE)mtext("X",side=1,line=1.5,outer=TRUE)

Probability Mass Function for Poisson Distribution



Without random coefficients, the standard Poisson model is:

$$\begin{aligned} &\text{begin\{equation\}} \\ &\log E(y_{\{i\}}) = \alpha + X'_{\{i\}} \beta \\ &\text{end\{equation\}} \end{aligned}$$

The log link is the canonical link function for the Poisson distribution, and the expected value of the response is modeled.

With random coefficients, for example a random intercept, the model becomes:

$$\begin{aligned} &\text{begin\{equation\}} \\ &\log E(y_{\{ij\}}|u_{\{j\}}) = \alpha + X'_{\{ij\}} \beta + u_{\{j\}} \\ &\text{end\{equation\}} \end{aligned}$$

Where $(y_{\{ij\}})$ is the observation for individual (i) in group (j) and $(u_{\{j\}})$ is the random effect for group (j) . Thus the two distributions are:

$$\begin{aligned} &\text{begin\{equation\}} \\ &y \sim \text{Pois}(\lambda) \end{aligned}$$

end{equation}

and

begin{equation}

$u \sim N(0, \sigma^{\{2\}})$

end{equation}

The random coefficient model is conditional on the random effect.

To show what this means, consider a simple model:

begin{equation}

$\log E(y_{\{ij\}}|u_{\{j\}}) = -.5 + .3x_{\{ij\}} + u_{\{j\}}$

end{equation}

In the original units, this becomes:

begin{equation}

$E(y_{\{ij\}}|u_{\{j\}}) = \exp(-.5 + .3x_{\{ij\}} + u_{\{j\}})$

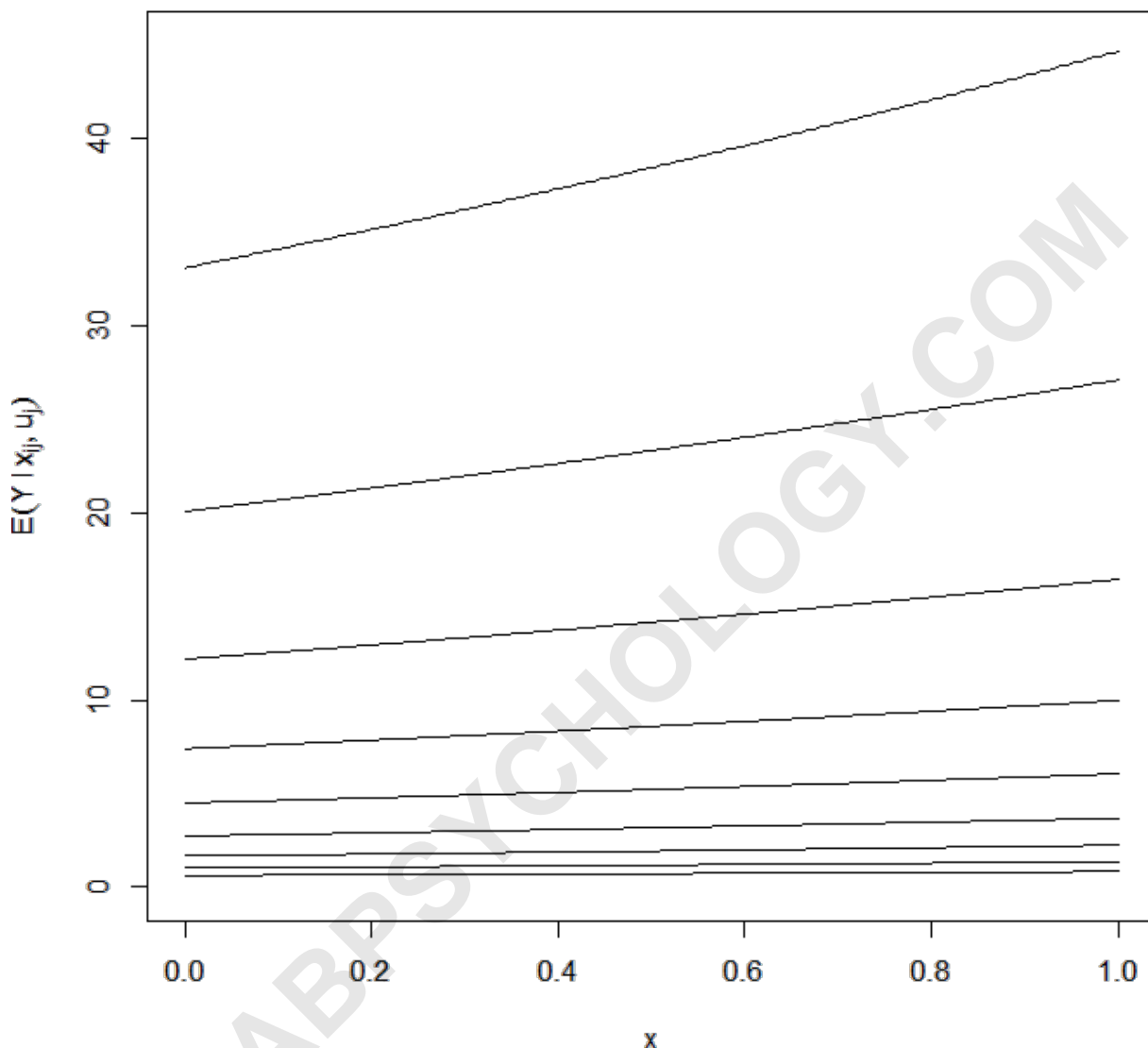
end{equation}

Now look what happens when we graph the estimated change for a 1 unit change in x for values of the random variable (u) ranging from 0 to 4 by increments of .5.

```
par(mfrow=c(1,1))
f<-function(x_ij,u_j){
  exp(-0.5+x_ij*0.3+u_j)}for(iinseq(0,4,0.5)){curve(f(x,u_j=i)
,from=0,to=1,n=200,add=i>0,ylim=c(0,45),ylab="")}title(yl
ab=bquote(E(Y~I~x,u)),main="Effect of x for different
random effects")
```

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Effect of x for different random effects



Clearly, on the scale of the original units, a 1 unit increase in x has different effects depending on the value of u , hence the conditionalness of the model. Population "average" effects can be

obtained by integrating out the random effect or by fitting a marginal model such as using GEEs.

Although the outcome is assumed to have a Poisson distribution, the random effect (in the above example, u) is typically assumed to have a Gaussian distribution.

Description of the data

The example data we use comes from a sample of the high school and beyond data set, with a made up variable, number of awards a student receives, awards. Our main predictor will be sex, female, and students are clustered (grouped) within schools, cid. First, we will load all the packages required for these examples.

```
## load foreign package to read Stata data files
require(foreign)
## ggplot2 package for graphs
require(ggplot2)
```

You need to install the `glmmADMB()` package via

instructions given in <http://glmmadmb.r-forge.r-project.org/>. Note that if you are not able to install the glmmADMB() package you may need to use Rtools and compile from source, use Linux, or ask the authors of the package for more information.

```
install.packages("R2admb")
```

```
install.packages("glmmADMB",  
repos="http://glmmadmb.r-forge.r-project.org/repos",  
type="source")
```

```
require(glmmADMB)
```

```
## load lme4 package  
require(lme4)
```

```
## read in data  
dat<-read.dta("https://stats.idre.ucla.edu/stat/data/hsbdemo.dta")  
dat$cid<-factor(dat$cid)
```

```
## look at the first few rows of the dataset  
head(dat)
```

```
id female ses schtyp prog read write math science  
socst honors awards cid
```

1 45 female low public vocation 34 35 41 29 26 not enrolled 0 1

2 108 male middle public general 34 33 41 36 36 not enrolled 0 1

3 15 male high public vocation 39 39 44 26 42 not enrolled 0 1

4 67 male low public vocation 37 37 42 33 32 not enrolled 0 1

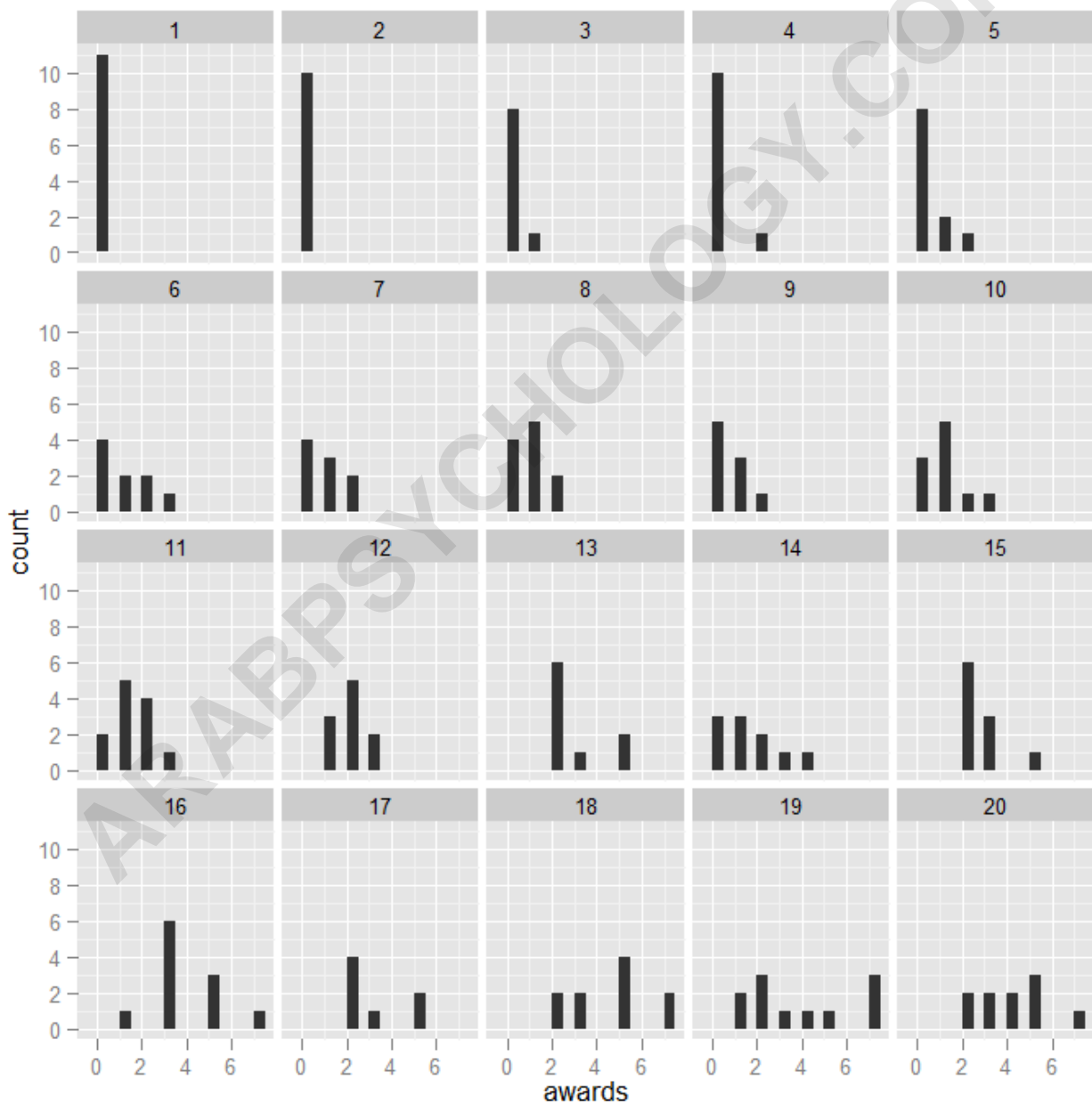
5 153 male middle public vocation 39 31 40 39 51 not enrolled 0 1

6 51 female high public general 42 36 42 31 39 not enrolled 0 1

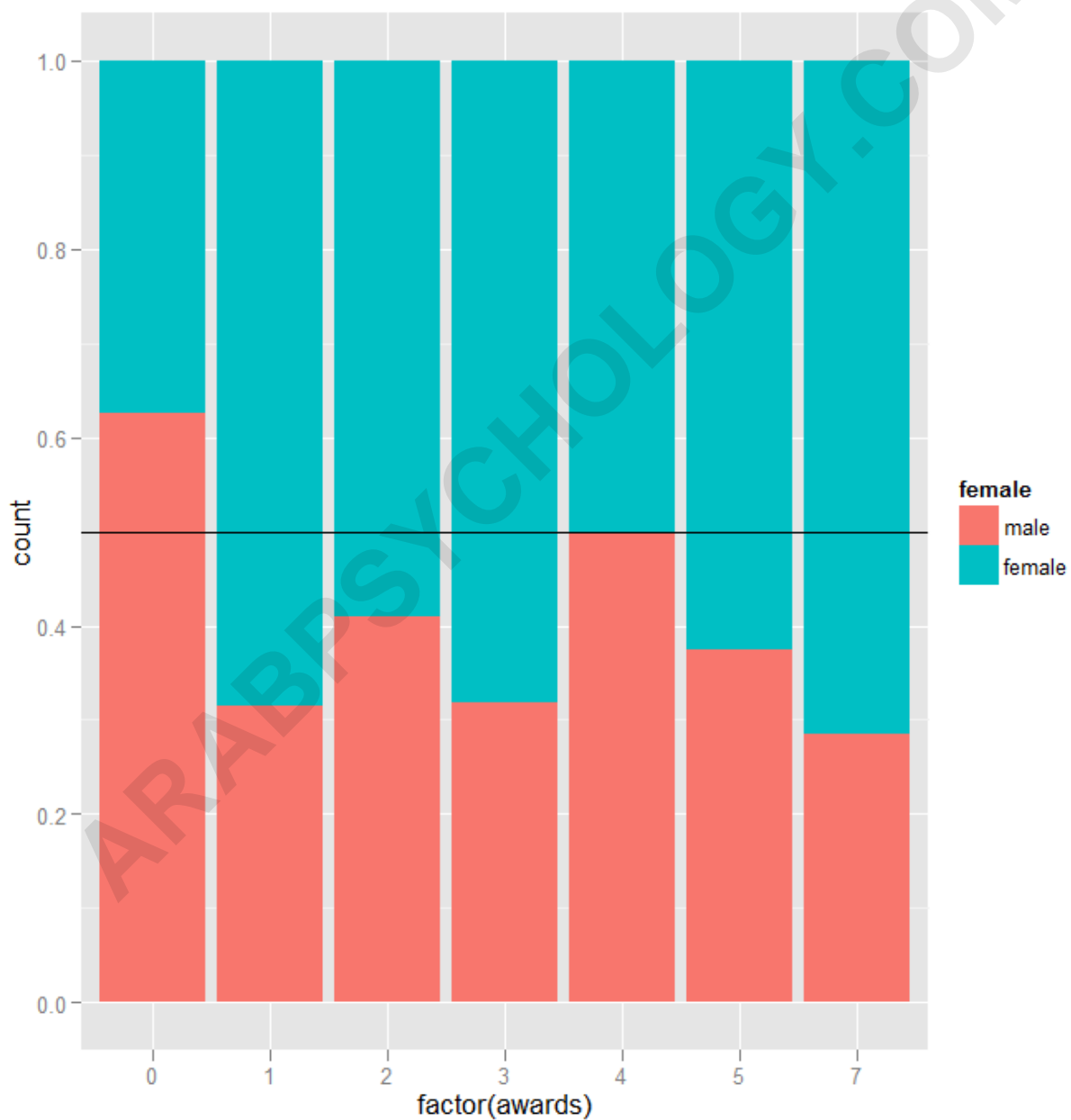
We can get a sense of the distributions in the data using the `ggplot2` package. The first plot is just histograms of number of awards for every cid. The second is a filled density plot. The density sums to 1 and the fill shows the distribution of female at every level of awards. If the distribution of female is equal across all awards, they would fall on the horizontal line.

awards by school

```
ggplot(dat,aes(awards))+geom_histogram(binwidth=0.5)  
+facet_wrap(~cid)
```



density of awards by sex, line at .5 is the null of no sex differences##
in number of
awardsggplot(dat,aes(factor(awards)))+geom_bar(aes(fill=female),position="fill")+ geom_hline(yintercept =0.5)



Analysis methods you might consider

Random coefficient poisson model analysis

Because generalized linear mixed models (GLMMs) such as random coefficient poisson models are rather difficult to fit, there tends to be some variability in parameter estimates between different programs. We will demonstrate the use of two packages in R that are able to fit these models, lme4 and glmmADMB.

```
## fit a random intercept only model using the Laplace
approximation## (equivalent to 1 point evaluated per
axis in Gauss-Hermite## approximation)m1a<-
glmer(awards~1+(1|cid),data=dat,family=poisson(link="l
og"))## fit a random intercept only model using 100
points per axis in the## adaptive Gauss-Hermite
approximation of the log likelihood more points##
improves accuracy but will take longer m1b<-
glmer(awards~1+(1|cid),data=dat,family=poisson(link="l
og"), nAGQ=100)## compare (only slightly
different) rbind(m1a=coef(summary(m1a)),m1b=coef(su
mmary(m1b)))
```

```
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -0.008853 0.2826 -0.03133 0.9750
## (Intercept) -0.009280 0.2834 -0.03275 0.9739
```

```
## view summary output from the more accurate
modelsummary(m1b)
```

```
Generalized linear mixed model fit by maximum
likelihood (Adaptive Gauss-Hermite Quadrature, nAGQ
=
```

```
100)
```

```
Family: poisson ( log )
```

```
Formula: awards ~ 1 + (1 | cid)
```

```
Data: dat
```

```
AIC BIC logLik deviance df.resid
228.6 235.2 -112.3 224.6 198
```

```
Scaled residuals:
```

```
Min 1Q Median 3Q Max
```

```
-1.3857 -0.5260 -0.3383 0.3379 3.3769
```

```
Random effects:
```

```
Groups Name Variance Std.Dev.
```

```
cid (Intercept) 1.458 1.207
```

Number of obs: 200, groups: cid, 20

Fixed effects:

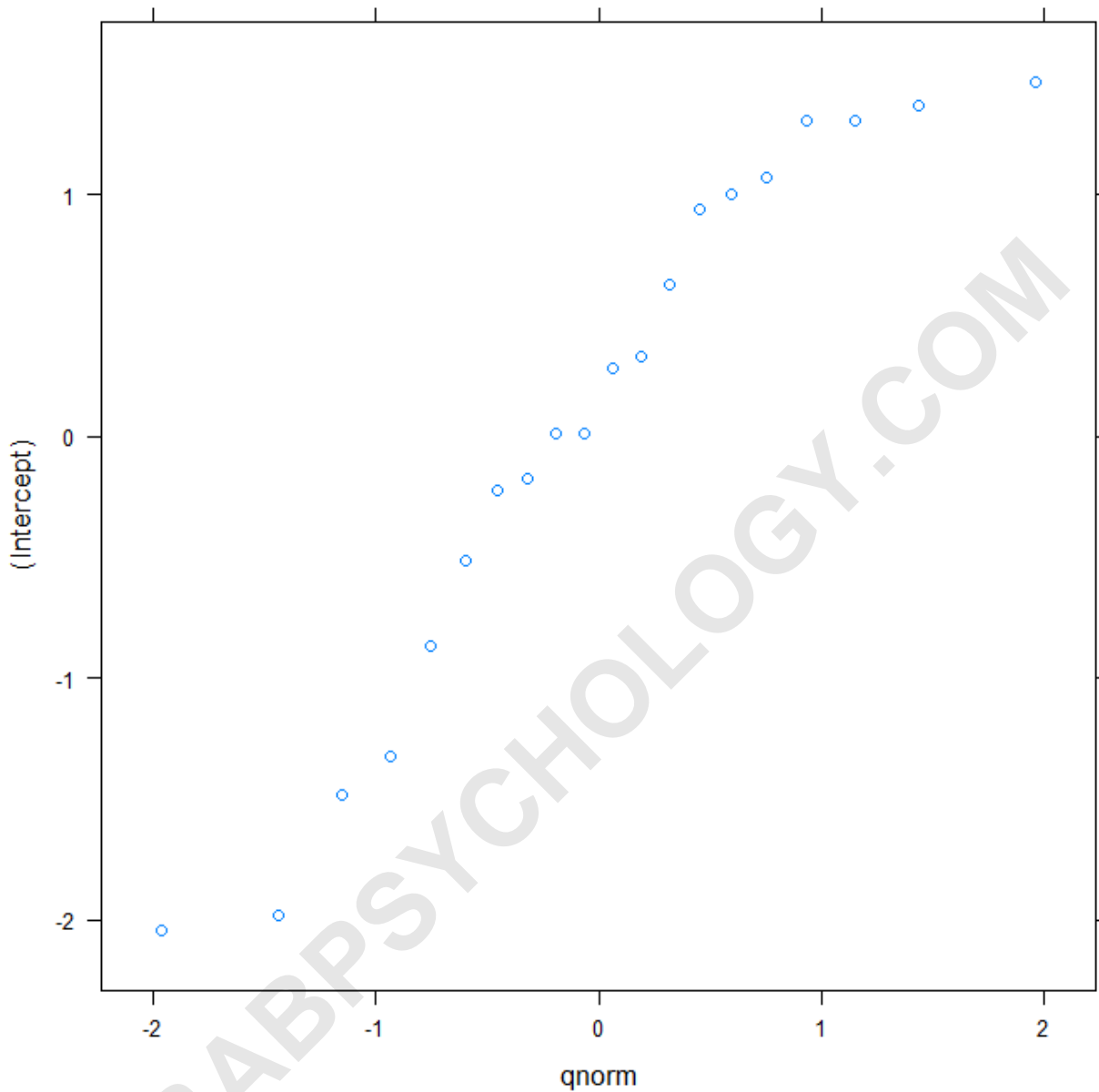
Estimate Std. Error z value Pr(>|z|)

(Intercept) -0.009575 0.290294 -0.033 0.974

We can visually inspect the random effects with a Q-Q plot or a caterpillar plot

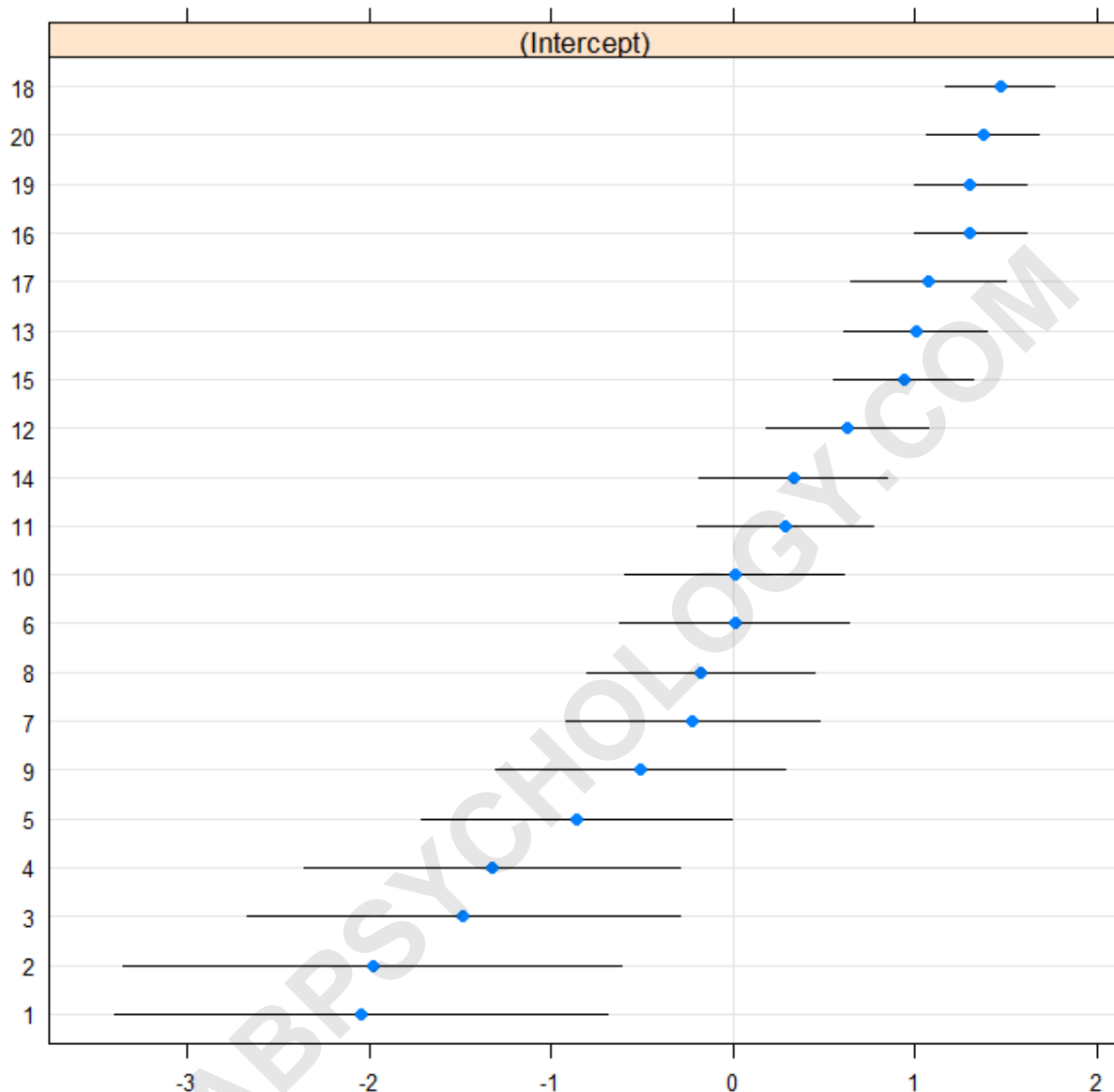
QQ plotplot(ranef(m1b))

\$cid



```
## Caterpillar  
plotlattice::dotplot(ranef(m1b,postVar=TRUE))
```

```
## $cid
```



The estimate for the intercept is essentially 0, although the random effects variance indicates that there is some variability in the intercepts between schools. Now we will add in female as

an explanatory variable.

m2 <-

```
glmer(awards~1+female+(1|cid),data=dat,family=poisson(link="log"),nAGQ=100)summary(m2)
```

Generalized linear mixed model fit by maximum likelihood (Adaptive Gauss-Hermite Quadrature, nAGQ

=

100)

Family: poisson (log)

Formula: awards ~ 1 + female + (1 | cid)

Data: dat

AIC BIC logLik deviance df.resid

221.1 231.0 -107.6 215.1 197

Scaled residuals:

Min 1Q Median 3Q Max

-1.5312 -0.5919 -0.3304 0.2047 2.8806

Random effects:

Groups Name Variance Std.Dev.

cid (Intercept) 1.431 1.196

Number of obs: 200, groups: cid, 20

Fixed effects:

Estimate Std. Error z value Pr(>|z|)

(Intercept) -0.2229 0.2975 -0.749 0.45370

femalefemale 0.3632 0.1193 3.044 0.00234 **

Signif. codes: 0 '*' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1**

Correlation of Fixed Effects:

(Intr)

femalefemal -0.252

There appears to be a fairly strong effect of females such that

females tend to get more awards than males. Now we will fit the same

models using the glmmADMB package

random intercept only model

library(glmmADMB)

m.alt1 <-

glmmadmb(awards~1+(1|cid),data=dat,family="poisson",link="log")

m.alt2 <-

```
glmmadmb(awards~1+female+(1|cid),data=dat,family="poisson",link="log")
```

```
summary(m.alt1)
```

Call:

```
glmmadmb(formula = awards ~ 1 + (1 | cid), data = dat, family = "poisson", link = "log")
```

AIC: 575.1

Coefficients:

Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-0.00881	0.28902	-0.03 0.98

Number of observations: total=200, cid=20

Random effect variance(s):

Group=cid

Variance	StdDev
(Intercept)	1.443 1.201

Log-likelihood: -285.528

```
summary(m.alt2)
```

Call:

```
glmmadmb(formula = awards ~ 1 + female + (1 | cid),  
data = dat,  
family = "poisson", link = "log")
```

AIC: 567.5

Coefficients:

```
Estimate Std. Error z value Pr(>|z|)  
(Intercept) -0.222 0.296 -0.75 0.4534  
femalefemale 0.363 0.119 3.04 0.0023 **
```

Signif. codes: 0 '*' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1**

Number of observations: total=200, cid=20

Random effect variance(s):

Group=cid

Variance StdDev

```
(Intercept) 1.417 1.19
```

Log-likelihood: -280.769

The results from glmmadmb match closely with those from glmer.

Things to consider

There have been reports (e.g., Zhang et al, 2009) of inconsistency and substantial alpha inflation between different statistical packages and estimation techniques (e.g., using penalized quasi-likelihood, Laplace approximation, Gauss-Hermite quadrature).

See also

Although not the primary focus of this page, because of these known difficulties with GLMMs, we also show the Stata and SAS code to run the same models using a variety of estimation techniques.

SAS

The dataset can be downloaded from here: [hsbdemo](#)

```
data dat;  
set "c:temphsbdemo.sas7bdat";  
run;  
  
proc sort data=dat;
```

```
by cid descending female;  
run;
```

```
PROC GLIMMIX data = dat method=LAPLACE noclprint;  
class cid;  
model awards = / solution dist=poisson;  
random intercept / subject=cid;  
run;
```

```
PROC GLIMMIX data = dat method=QUAD(QPOINTS =  
100) noclprint;  
class cid;  
model awards = / solution dist=poisson;  
random intercept / subject=cid;  
run;
```

```
PROC GLIMMIX data = dat method=LAPLACE noclprint  
order = data;  
class female cid;  
model awards = female / solution dist=poisson;  
random intercept / subject=cid;  
run;
```

```
PROC GLIMMIX data = dat method=QUAD(QPOINTS =  
100) noclprint order = data;
```

```
class female cid;  
model awards = female / solution dist=poisson;  
random intercept / subject=cid;  
run;
```

*/*We can also manually parameterize the model in NL MIXED */*

```
PROC NL MIXED data = dat METHOD=GAUSS  
QPOINTS=100;  
eta = alpha + u;  
lambda = exp(eta);  
model awards ~ poisson(lambda);  
random u ~ normal(0, exp(2*log_sigma)) subject=cid;  
estimate 'intercept variance' exp(2*log_sigma);  
run;
```

```
PROC NL MIXED data = dat METHOD=GAUSS  
QPOINTS=100;  
eta = alpha + beta*female + u;  
lambda = exp(eta);  
model awards ~ poisson(lambda);  
random u ~ normal(0, exp(2*log_sigma)) subject=cid;  
estimate 'intercept variance' exp(2*log_sigma);  
run;
```

Random intercept only model, Laplace approximation)

Covariance Parameter Estimates			
Cov Parm	Subject	Estimate	Standard Error
Intercept	CID	1.4429	0.5970

Solutions for Fixed Effects					
Effect	Estimate	Standard Error	DF	t Value	Pr > t
Intercept	-0.00877	0.2890	19	-0.03	0.9761

Random intercept only model, adaptive Gauss-Hermite approximation with 100 evaluation points)

Covariance Parameter Estimates			
Cov Parm	Subject	Estimate	Standard Error
Intercept	CID	1.4579	0.6057

Solutions for Fixed Effects					
Effect	Estimate	Standard Error	DF	t Value	Pr > t
Intercept	-0.00954	0.2903	19	-0.03	0.9741

Random intercept plus female model, adaptive Gauss-Hermite approximation with 100 evaluation points)

Covariance Parameter Estimates			
Cov Parm	Subject	Estimate	Standard Error
Intercept	CID	1.4168	0.5867

Solutions for Fixed Effects						
Effect	FEMALE	Estimate	Standard Error	DF	t Value	Pr > t
Intercept	-0.2222	0.2963	19	-0.75	0.4625	
FEMALE	1	0.3632	0.1193	179	3.04	0.0027
FEMALE	0	0

Random intercept only model, adaptive Gauss-Hermite approximation with 100 evaluation points) using PROC NL MIXED

Parameter Estimates	Estimate	Standard Error	DF	t Value	Pr > t	95% Confidence Limits	Gradient	
alpha	-0.00957	0.2903	19	-0.03	0.9740	-0.6172	0.5980	2.881E-6
log_sigma	0.1885	0.2077	19	0.91	0.3756	-0.2463	0.6233	7.13E-8

Additional Estimates								
Label	Estimate	Standard Error	DF	t Value	Pr > t	Alpha	Lower	Upper
intercept variance	1.4579	0.6057	19	2.41	0.0264	0.05	0.1901	2.7257

Random intercept plus female model, adaptive Gauss-Hermite approximation with 100 evaluation points) using PROC NL MIXED

Parameter Estimates								
Parameter	Estimate	Standard Error	DF	t Value	Pr > t	95% Confidence Limits	Gradient	
alpha	-0.2229	0.2975	19	-0.75	0.4629	-0.8456	0.3998	8.17E-8
beta	0.3632	0.1193	19	3.04	0.0067	0.1134	0.6129	-0.00004
log_sigma	0.1793	0.2079	19	0.86	0.3991	-0.2558	0.6145	9.151E-6

Additional Estimates								
Label	Estimate	Standard Error	DF	t Value	Pr > t	Alpha	Lower	Upper
intercept variance	1.4314	0.5952	19	2.40	0.0265	0.05	0.1856	2.6773

Stata

use <https://stats.idre.ucla.edu/stat/data/hsbdemo>, clear

***Random intercept model, Laplace**

xtmepoisson awards || cid:, laplace var

***Random intercept model, adaptive Gauss-Hermite, 100 points**

xtmepoisson awards || cid:, intpoints(100) var

***Random intercept and female**

xtmepoisson awards female || cid:, laplace var

***Random intercept and female**

xtmepoisson awards female || cid:, intpoints(100) var

Random intercept only model, Laplace approximation

awards | Coef. Std. Err. z P>|z|

-----+-----
_cons | -.0088 .2890208 -0.03 0.976 -.5752703 .5576703

Random-effects Parameters | Estimate Std. Err.

-----+-----
cid: Identity |

var(_cons) | 1.442954 .5970273 .6413056 3.246685

Random intercept only model, adaptive Gauss-Hermite approximation with 100 evaluation points)

awards | Coef. Std. Err. z P>|z|

-----+-----
_cons | -.0095717 .2902932 -0.03 0.974 -.5785359
.5593924

Random-effects Parameters | Estimate Std. Err.

cid: Identity |

var(_cons) | 1.45791 .6057226 .645772 3.291411

Random intercept plus female model, Laplace approximation

awards | Coef. Std. Err. z P>|z|

female | .3632063 .1193049 3.04 0.002 .129373 .5970396

_cons | -.2221984 .2963088 -0.75 0.453 -.8029531
.3585562

Random-effects Parameters | Estimate Std. Err.

cid: Identity |

var(_cons) | 1.416836 .5867228 .6292602 3.190133

Random intercept plus female model, adaptive Gauss-Hermite approximation with 100 evaluation points)

awards | Coef. Std. Err. z P>|z|

female | .3631578 .1193071 3.04 0.002 .1293202 .5969954
_cons | -.2229269 .29752 -0.75 0.454 -.8060553 .3602015

Random-effects Parameters | Estimate Std. Err.

cid: Identity |
var(_cons) | 1.431444 .5952188 .6336221 3.23384

References